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ESTIMATING THE CONTRIBUTION OF EXPERIMENTAL TESTS TO THE ARMED SERVICES VOCATIONAL APTITUDE BATTERY

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13. ABSTRACT (Maximum 200 words)

For the only way to accurately determine how much an experimental test will add to the Armed Services Vocational Aptitude Battery (ASVAB) in predicting a subsequent criterion would be to administer the experimental test along with ASVAB in the operational setting. This simply is not feasible to do on a routine basis. Ordinarily, experimental tests are administered to individuals already in the service and an attempt is made to determine how much predictive efficiency these tests add to ASVAB subtests administered some months previously. This approach has been justly criticized because abilities could have changed between the time of ASVAB administration and experimental test administration. Because of the availability of ASVAB test-retest data on over 4,000 subjects, it was possible in the present research effort to address the ability changes in terms of a meta-analysis of differences between test and retest validities for grades in 69 technical school courses. It was hypothesized that the retest scores should have higher validities to the extent that abilities had changed. This follows from the fact that retest scores measure abilities at the time of course entry.

With the exception of slightly higher retest validities for ASVAB measures of technical knowledge, the test and retest validities centered at about the same level. It was concluded that reasonable

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trust can be placed in estimates of the contributions of experimental tests to ASVAB in the prediction of subsequent criteria, even though the ASVAB scores are collected some months prior to administration of the experimental tests. ($\leq r_{\odot}$)

Another approach to estimating the contribution of experimental tests to ASVAB would be to readminister ASVAB concurrently with the experimental tests. The present research indicates that this approach would, on the average, yield values comparable to those which would be obtained if the experimental tests had been administered at the operational testing stations, along with ASVAB.

Validities of experimental tests for concurrent criterion measures administered in the same testing session appear to be moderately inflated. This inflation is hypothesized to be due in part to a few non-cooperative subjects who operate at a reduced level of effort on all tests when they are told that scores will not affect their careers. It is recommended that, to the extent possible, such non-cooperative subjects be identified and removed from the validation sample. Estimated contributions may still be slightly inflated, but results from the present study suggest that such inflation is not likely to be large.



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ESTIMATING THE CONTRIBUTION OF EXPERIMENTAL TESTS TO THE ARMED SERVICES VOCATIONAL APTITUDE BATTERY

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This publication is primarily a working paper. It is published solely to document work performed.

SUMMARY

This paper is concerned with how to estimate the contribution of experimental tests to the Armed Services Vocational Aptitude Battery (ASVAB) in accounting for criterion variance. Ordinarily, experimental tests are administered to individuals already in the service and an attempt is made to determine how much predictive efficiency these tests contribute to ASVAB subtests, which were administered some months previously. This approach has been justly criticized, in that ability levels could have changed between the time of ASVAB administration and experimental test administration. In the LAMP program, more serious problems are faced. Often an attempt is made to determine how much power the LAMP tests add to ASVAB when the criterion measure is administered concurrently with the experimental tests. Thus, important questions need to be answered. How much do the abilities of Air Force applicants change between the time of operational testing and entry into service? Would scores from ability tests administered at time of entry into technical schools have higher validities for course grades than would scores collected some months earlier in the operational testing program? How much are estimates of the validity added by LAMP tests to ASVAB inflated because of concurrency effects? It was possible to address these and other questions in the present research because of the availability of ASVAB test-retest data on over 4,000 cases. Investigation of ability changes between test and retest was made by meta-analyses of validities for technical school grades. It was hypothesized that retest scores should have higher validities, since they cap abilities at the time of school entry. Concurrency effects were analyzed by internal analysis of ASVAB test-retest data.

With the exception of slightly higher retest validities for ASVAB measures of technical knowledge, the test and retest validities centered at about the same level. The validities of ASVAB retest scores may have been slightly underestimated because of situational variance during the experimental retest situation; however, this situational variance will lead to underestimates of the validity of any experimental test for subsequent operational criteria. It was concluded that, until better information becomes available, the estimated contribution of experimental tests to ASVAB in the prediction of subsequent criteria be accepted at face value.

Another approach to estimating the contribution of experimental tests to ASVAB would be to readministered ASVAB concurrently with the experimental tests. The present research indicates that this approach would, on the average, yield values comparable to those which would be obtained if the experimental tests had been administered at the operational testing stations with ASVAB.

Validities of experimental tests for concurrent criterion measures (administered in the same testing session) appear to be moderately inflated. This inflation is hypothesized to be due in part to a few non-cooperative subjects who operate at a reduced level of effort on all tests when they are told that scores will have no effect on their careers. It is recommended that, to the extent possible, such non-cooperative subjects be identified and removed from validation samples. Estimated contributions may still be slighly inflated, but results from the present research suggest that such inflation is not likely to be large.

PREFACE

Support for this effort was provided by the Air Force Human Resources Laboratory (AFHRL), under contract F41689-86-D-0052. The data analyzed and reported in this paper were originally collected by Maj. John Welsh and Capt Randy Massey for use in the ASVAB quality control program. I am indebted to Mr. Calvin Fresne and Mrs. Janice Herford for assistance in locating data files and reorganizing them for analysis. Ms Carmen Peña provided assistance in proofing tables and locating references. Appreciation is expressed to Dr. William Alley, Dr. Patrick Kyllonen, Dr. William Tirre, Dr. Scott Chaiken, and Dr. Sidney Irvine for their suggestions and encouragement. I also wish to acknowledge the assistance of Dr. Malcom Ree and Mr. William Phalen for comments concerning analysis alternatives. Finally, I am deeply indebted to Dr. Lloyd Humphreys and Dr. Daniel Woltz, both of whom read early drafts of this paper and made many helpful suggestions leading to its improvement.

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Estimating the Contribution of Experimental Tests to the Armed Services Vocational Aptitude Battery

I. INTRODUCTION

Advances in cognitive theory and the general availability of microcomputers have stimulated research during the last decade on the development of computer-based tests designed to measure individual differences in regnition. A question currently being addressed is whether these new theory-based tests will contribute anything to conventional paper-and-pencil ability tests in predicting subsequent learning and performance criteria. During the last few years, scientists working in the Air Force Learning Abilities Measurement Program (LAMP)¹ have conducted research involving validation of new experimental tests against concurrent or subsequent learning and performance measures. A normal practice is to obtain official scores from the previously administered Armed Services Vocational Aptitude Battery (ASVAB) and compute how much the experimental tests contribute to that battery in accounting for criterion variance. Results from this procedure relating to the incremental validity of the experimental tests may be inflated. The ASVAB validities may be attenuated because individuals have changed in the time between ASVAB and criterion test administrations. Because the new experimental tests are administered at or near the time of the criterion measures, they are not subject to this attenuation.

Humphreys and others (Humphreys, 1960; Humphreys & Davey, 1988; Humphreys, Davey, & Park, 1985; Humphreys, Parsons, & Park, 1979) have shown that when test scores or school grades are collected on the same individuals at several points in time, the intercorrelations of scores or grades yield a quasi-simplex matrix. I'hat is, high correlations are obtained between successive scores or grades across a short time period; but the longer the time intervening between data collection, the lower the correlations. However, it should be noted that these studies dealt with children or with young adults in college. In both instances significant changes in knowledge acrosstime should have been expected.

It may be that the knowledge changes in Air Force selectees between testing and service entry are far less dramatic than those demonstrated by Humphreys. Over 40% of the individuals tested in LAMP enter the Air Force within 6 months of operational testing, and roughly 87% come in within 12 months of operational testing. During this intervening time, many of these individuals are not in school at all, although some are tested in high school and take additional courses before graduation and entry into the Air Force.

The only way to accurately determine how much an experimental test will contribute to ASVAB in predicting a subsequent criterion would be to ad-

¹ LAMP is a basic research program which is jointly sponsored by the Air Force Office of Scientific Research (AFOSR) and the Air Force Human Resources Laboratory (AFHRL). LAMP seeks to understand how individuals process information in order to perceive, store, remember, solve problems, and acquire knowledge and skills. An ultimate goal is to develop new ability measures that can be used in the Air Force Personnel Selection and Classification Program.

minister it along with the ASVAB in the operational setting. This is simply not feasible to do on a routine basis at the present time.

A second approach would be to readminister the ASVAB along with experimental tests after subjects have entered the Air Force. Then one could find out how much power the experimental tests add to ASVAB in predicting concurrent or subsequent criteria. This would eliminate the time differential between ASVAB and experimental test administrations, but it might not accurately show what the validities and joint validities would have been had these measures been administered in the operational testing stations some months previously. Furthermore, insufficient testing time is available for this approach.

The third approach is the one currently being used in LAMP research; that is, to determine how much the scores on experimental tests, administered at or near the time of criterion data collection, contribute to operational ASVAB scores in accounting for variance in the criterion. This is undoubtedly the weakest approach, but the only one currently feasible. No information is yet available as to how fallible such estimates might be.

Thus, serious questions must be addressed regarding the LAMP validation process. How much do the abilities of Air Force applicants change between the time of operational testing and the time of entry into the service? Would scores from ability tests administered at time of entry have higher validities for technical school grades than would scores collected some months earlier in the operational testing program? How much are estimates of the validity contributed by LAMP tests inflated because of concurrency effects? The present study was designed to provide information bearing on such questions.

II. APPROACH

During the 1984-1987 time period, the Air Force readministered the ASVAB to random samples of airmen on the 6th day of their Basic Military Training. This is the same day that experimental LAMP data are collected. The present investigation consists of analyses of these ASVAB test-retest data in an effort to address the predictive-concurrent validity question.

III. METHOD

ASVAB test-retest scores were available on all subtests for 4,077 cases. The sample was divided into three subsamples: (a) those who entered the Air Force within 6 months after initial testing (N = 1,774), (b) those who entered between 6 and 12 months after testing (N = 1,785) and (c) those who entered 12 or more months after testing (N = 518).

Three types of analyses were conducted. First, test-retest correlations, means, and standard deviations were computed for all 10 ASVAB subtests. These data provide a rough indication of the stability of test scores across time. Second, test and retest data were correlated with technical school grades for a variety of courses. Due to concurrency effects, one would expect the retest scores to have higher validities. Finally, each of eight ASVAB subtests at Time 2 (retest) was selected as a "criterion," and in each instance. a second related test was selected as an "experimental" predictor. It was then possible to compare how much the experimental predictor added to "ASVAB" (the remaining eight subtests) in predicting the criterion under two conditions: (a) The ASVAB was administered months earlier under operational conditions, and (b) the ASVAB was readministered concurrently with the experimental and criterion variables. Each of these analyses is discussed below.

IV. RESULTS

Test-Retest Correlations, Means, and Standard Deviations

Table 1 presents the names of the various subtests in ASVAB, along with the abbreviations which are used throughout this paper. Table 2 reports test-retest correlations for the ASVAB subtests, separately for each of the three subsamples and the total sample. This table also reports median internal consistency reliability estimates for eight of the subtests, as reported by Ree, Mullins, Mathews. & Massey (1982). Table 3 reports

Table 1
ASVAB TESTS AND ABBREVIATIONS

| TEST | ABBR. |
|--------------------------|-------|
| General Science | GS |
| Arithmetic Reasoning | AR |
| Word Knowledge | WK |
| Paragraph Comprehension | PC |
| Numerical Operations | NO |
| Coding Speed | CS |
| Auto-Shop Information | AS |
| Mathematics Knowledge | MK |
| Mechanical Comprehension | MC |
| Electrical Information | ΕI |
| | |

Table 2
ASVAB TEST-RETEST CORRELATIONS

| TEST | T 0-6MOS | 6-12HOS | >12MOS | TOTAL | r11 |
|-----------|----------|---------|--------|-------|------|
| M | 1174 | 1785 | 518 | 4077 | |
| G\$ | .744 | .736 | .737 | .740 | .84 |
| AR | .768 | .762 | .773 | .766 | .90 |
| LK | .774 | . 765 | . 798 | .772 | .92 |
| PC | .493 | .407 | .350 | .439 | .80 |
| NO | .690 | .672 | .620 | .674 | * |
| cs | -694 | .651 | .596 | .663 | • |
| AS | .800 | .808 | .776 | .800 | . 88 |
| MK | .839 | .819 | .617 | .828 | .87 |
| MC | .716 | .704 | .714 | .710 | . 86 |
| ΕĮ | .692 | .687 | .662 | .685 | . 83 |
| | | | | | |

Notes: r11 refers to internal consistency reliabilities. * = not computed.

test-retest means and standard deviations for the various subsamples and total sample.

Data in Table 2 reveal some lack of stability in scores across time, especially for the Paragraph Comprehension, Coding Speed, Numerical Operations, and Electrical Information subtests. Part of this is attributable to a lack of internal consistency in the tests themselves. Paragraph Comprehension and Electrical Information have the two lowest reliability coefficients of those considered. Furthermore, Numerical Operations and Coding Speed are both speeded tests, and speeded tests are noted for their lack of stability across time. The test-retest correlations are attenuated somewhat by the use of several test forms. However, great pains have been taken to produce ASVAB forms which are equivalent in terms of item characteristics, and such attenuation is judged to be of minor consequence.

Data for the three time periods reported in Table 2 do not form a simplex. Although the correlations for the 6-12 months group are slightly lower that those for the 0-6 months group, several of the correlations for the over 12 months group are actually higher than those for the 0-6 months group. The most striking feature in Table 2 is the consistency of values across time. Of course, these data do not represent correlations on the same individuals across time. They are from independent subsamples. The small differences noted could be due entirely to sampling fluctuations.

Table 3 reports test-retest means and standard deviations for subtests, separately for each of the subsamples and the total sample. The test-retest means show a remarkable stability across time. Although some minor fluctuations can be observed from subtest to subtest, it appears that, on the whole, forgetting, learning, practice effects, regression effects and situational variance at Time 2 are fairly well balanced out. The standard deviations appear to be generally higher at Time 2, a matter which will be discussed later in this paper.

Although the data in Tables 2 and 3 suggest little difference between test and retest scores regardless of interval times, firm conclusions concerning

TABLE 3

ASVAB TEST-RETEST MEANS AND STANDARD DEVIATIONS

MEANS

| 0-6 MONTHS (N=1774) | | | | 6-12 MONTHS (N=1785) | | ONTHS | TOTAL SAMPLE (N=4077) | | |
|------------------------|-------|-------|-------|-------------------------|-------|-------|-----------------------|-------|--|
| TEST | TIME1 | TIME2 | TIME1 | TIME2 | TIME1 | TIME2 | - | TIME2 | |
| GS | 18.7 | 18.4 | 18.6 | 18.6 | 18.9 | 19.1 | 18.7 | 18.6 | |
| AR | 22.9 | 22.1 | 22.8 | 21.9 | 23.3 | 22.6 | 22.9 | 22.1 | |
| WK | 29.2 | 29.4 | 28.8 | 29.5 | 29.2 | 29.9 | 29.1 | 29.5 | |
| PC | 12.5 | 12.4 | 12.4 | 12.3 | 12.4 | 12.5 | 12.4 | 12.4 | |
| NO | 40.8 | 40.6 | 40.9 | 40.3 | 41.6 | 41.1 | 40.9 | 40.5 | |
| cs | 53.3 | 54.3 | 52.3 | 53.4 | 53.0 | 53.0 | 52.8 | 53.7 | |
| AS | 18.4 | 18.7 | 18.2 | 18.5 | 17.9 | 18.7 | 18.3 | 18.6 | |
| MK | 16.6 | 16.2 | 16.9 | 16.3 | 17.2 | 16.8 | 16.8 | 16.3 | |
| MC | 17.9 | 18.0 | 17.6 | 17.8 | 17.7 | 18.3 | 17.7 | 18.0 | |
| ٤١ | 14.1 | 13.9 | 13.9 | 13.9 | 13.8 | 14.0 | 14.0 | 13.9 | |

STANDARD DEVIATIONS

| | 0-6 MONTHS | 6-12 MONTHS | >12 MONTHS | TOTAL SAMPLE |
|------|-------------|------------------|-------------|--------------|
| | (N=1774) | (N=1785) | (N=518) | (N=4077) |
| TEST | TIME1 TIME2 | TIME1 TIME2 | TIME1 TIME2 | TIME1 TIME2 |
| GS | 3.66 3.77 | 3.53 3.68 | 3.29 3.56 | 3.56 3.71 |
| AR | 4.64 5.35 | 4.59 5.17 | 4.52 4.93 | 4.60 5.22 |
| WK | 4.37 4.29 | 4.34 4.11 | 4.10 3.90 | 4.33 4.16 |
| PC | 1.93 1.99 | 1.92 1.92 | 1.78 1.92 | 1.91 1.95 |
| NO | 7.24 7.67 | 7.01 7.58 | 7.01 7.33 | 7.12 7.59 |
| CS | 11.94 12.30 | 11.41 12.23 | 11.60 11.48 | 11.67 12.18 |
| AS | 4.39 4.27 | 4.33 4.41 | 4.41 4.36 | 4.37 4.34 |
| MK | 5.04 5.23 | 4.74 5.13 | 4.68 5.02 | 4.87 5.16 |
| MC | 3.98 3.98 | 4.08 3.98 | 4.00 3.89 | 4.03 3.97 |
| ΕI | 3.15 3.37 | 3.12 3.36 | 3.19 3.44 | 3.14 3.38 |

ability changes across time cannot be drawn from comparative analyses of data from three different subsamples. Therefore, emphasis was shifted to the second approach, which involved validation of Time 1 (test) and Time 2 (retest) scores against technical school course grades.

Test-Retest Validation Against Technical School Grades

The ASVAB subtest test-retest correlations are low enough to allow for significant changes to have occurred in the abilities of applicants between the time of operational testing and the individuals' entry into the Air Force. To the extent that such changes did occur, one would expect the retest scores to have higher vatidities for technical school grades. This follows from the fact that the scores gathered just prior to entry into the technical school should be the best indicators of abilities at that time. Two approaches were taken to evaluate this hypothesis. One involved using a simple signs test of the differences in validity coefficients yielded by the test and retest data. The second involved more sophisticated meta-analyses.

Signs Tests. ASVAB scores for the 4,077 subjects included in the study were matched with operational files to obtain technical school grades. Table 4 reports test and retest validities of the 10 ASVAB subtests for grades in each of the 20 courses in which the matching process yielded 40 or more subjects. Thus, 200 pairs of validity coefficients were available to test the hypothesis. Because test and retest scores came from the same individuals, 200 t-values were computed using Hotelling's formula for correlated data, one for each test-retest validity pair. These t-values are reported in Table 5, ordered by ASVAB subtest. A negative t-value indicates that a particular retest validity coefficient is higher than its associated test validity coefficient.

A cursory inspection of the data in Table 5 reveals a weak tendency for the retest validities to be higher. There are 110 negative t-values and 90 positive ones. Nevertheless, applying a simple signs test produces a Chi-Square of only -2.225, which is non-significant at the .05 level. The over-

all hypothesis must be rejected. It appears that, in general, retest validities are not significantly higher than test validities. Although test-retest validities did not differ across all subtests, it was hypothesized that ability changes were more likely to occur among the ASVAB subtests measuring technical knowledge, than among those measuring general ability or perceptual speed. Certainly technical knowledge is subject to learning and forgetting across time. Therefore, the ASVAB subtests were sorted into three categories for further evaluation, as follows:

TECHNICAL KNOWLEDGE Electrical Information Mechanical Comprehension General Science Auto-Shop Information

GENERAL ABILITY
Arithmetic Reasoning
Word Knowledge
Mathematics Knowledge
Paragraph Comprehension

PERCEPTUAL SPEED Numerical Operations Coding Speed

Data in Table 5 reveal that the four technical knowledge subtests yielded 52 negative t-values but only 28 positive t-values. In this instance, the simple signs test produces a Chi-Square of -7.212, which is significant beyond the .01 level. Thus, the hypothesis of higher retest validities for ASVAB measures of technical knowledge is supported. The signs tests for differences in the general ability and perceptual speed ASVAB measures produced Chi-Squares of 0.812 and -0.360, respectively, neither of which is significant at the .05 level.

Meta-Analyses. As a second, and much more powerful approach, a series of meta-analyses were conducted for the three categories of tests mentioned above (technical knowledge, general ability, and perceptual speed), and for all subtests combined. Results of these analyses are presented

Table 4
TEST-RETEST CORRELATIONS WITH SCHOOL GRADES

| 04470 | | | | | | | | | | | |
|--------|--------|---------------|----------|-----------|---------|--------|------------|-------|---------------|-------------|-------|
| 81130 | - | | RITY SPE | | | | | | | | |
| TEST | N | 22 2 2 2 2 | AR | WK . | PC | NO | CS | AS | MK 0.7// | NC 0.403 | EI |
| TIMET | 268 | | | 0.459 | | 0.110 | | 0.273 | 0.364 | 0.182 | 0.284 |
| TIME2 | 268 | 0.482 | 0.392 | 0.454 | 0.336 | 0.170 | 0.1/4 | 0.281 | 0.346 | C.130 | 0.353 |
| | | | | | | | | | | | |
| 43151 | • | | ICAL AIR | | | | | • | | | |
| TEST | X | GS | AR | WK. | PC | NO | CS | AS | HK | MC | ΕI |
| TIME1 | 223 | 0.435 | | | | | 0.112 | | | | |
| TIME2 | 223 | 0.394 | 0.416 | 0.315 | 0.316 | 0.106 | 0.129 | 0.389 | 0.419 | 0.358 | 0.479 |
| | | | | | | | | | | | |
| 81150 | • | | ECIALIST | • | | _ | | | | | |
| TEST | N | GS | AR | UK | PC | NO | CS | AS | MK | MC | EI |
| TIMET | 182 | 0.448 | 0.271 | 0.345 | 0.266 | 0.223 | 0.088 | 0.255 | 0.334 | 0.326 | 0.149 |
| TIMEZ | 182 | 0.468 | 0.336 | 0.337 | 0.271 | 0.274 | 0.169 | 0.268 | 0.324 | 0.275 | 0.257 |
| | | | <u>-</u> | | | | | | | | |
| 702308 | • | | INISTRAT | | | | | | | | |
| TEST | K | GS | AR | WK . | PC | NO | CS | AS | MK | MC | EI |
| TIME1 | 118 | 0.354 | | | | -0.024 | | 0.257 | 3 .367 | 0.392 | 0.332 |
| TIMEZ | 118 | 0.369 | 0.433 | 0.215 | 0.439 | 0.091 | 0.142 | 0.258 | 0.481 | 0.267 | 0.251 |
| | | | | | | | | | | | |
| 64530 | • | | ITORY MA | | | | | | | | |
| TEST | N | G\$ | AR | UK | PC | NO | , CS | AS | MK | MC | ΕI |
| TIME | 114 | 0.264 | 0.451 | 0.329 | | | 0.016 | | 0.273 | 0.059 | 0.078 |
| TIME2 | 114 | 0.250 | 0.345 | 0.274 | 0.181 | 0.084 | -0.066 | 0.202 | 0.323 | 0.186 | 0.171 |
| | | | | | | | | | | | |
| 81132 | (APPRN | | NFORCEN | ENT SPE | CIALIST | - | | | | | |
| TEST | N | GS | AR | WK | PC | NO | c s | AS | MK | MC | ΕI |
| TIMET | 110 | | | | | 0.126 | | 0.325 | | 0.372 | 0.310 |
| TIME2 | 110 | 0.542 | 0.406 | 0.401 | 0.362 | 0.209 | 0.167 | 0.418 | 0.346 | 0.340 | 0.455 |
| | | | | | | | | | | | |
| 57130 | - | | PROTECT | | | • | | | | | |
| TEST | h | GS | AR | ¥K | PC | NO | CS. | AS | MK | MC | EI . |
| TIME1 | 89 | | 0.302 | | | | 0.147 | | | 0.323 | 0.279 |
| TIMES | 89 | 0.417 | 0.3867 | 0.243 | 0.167 | 0.134 | 0.063 | 0.226 | 0.233 | 0.451 | 0.333 |
| 0007.1 | | Lemm n | | | | | | | | | |
| 90230 | • | | CAL SERV | | | • | | | | | |
| TEST | N | G\$ | AR | WK | PC | NO | C\$ | AS | MK | MC | 13 |
| TIMET | 86 | 0.482 | 9.350 | 0.379 | | | -0.076 | | 0.323 | 0.206 | 0.293 |
| TIMEZ | 86 | 0.473 | 0.557 | 0.462 | 0.286 | -0.077 | 0.018 | 0.508 | 0.228 | 0.330 | 0.465 |
| /3/30 | | | | | | | | | | | |
| ა3130 | • | | SPECIAL | | | | | | | | |
| TEST | N | G\$ | AR | WK | PC | NO | C\$ | AS | MK | MC | EI |
| TIME1 | | | 0.432 | | | | | | | | |
| TIME2 | 84 | 0.071 | 0.401 | 0.093 | 0.240 | 0.242 | 0.269 | 0.083 | 0.414 | 0.173 | 0.101 |
| | | | | | | | | | | | |
| 73230 | • | | DANEL SP | | • | | | | | | |
| TEST | W | G S | AR | WK | PC | NO | CS | AS | MK | MC | EI |
| TIME1 | 74 | | | 0.513 | | | -0.054 | | | | |
| TIME2 | 74 | 0.241 | 0.475 | 0.441 | 0.379 | 0.112 | 0.151 | 0.123 | 0.376 | 0.231 | 0.267 |

Table 4 (Cont.)

TEST-RETEST CORRELATIONS WITH SCHOOL GRADES

| /2472 | /4000W | IET S | ENGINE M | IE CUANT | • • | | | | | | |
|--------|--------|------------|----------|------------|-----------|----------|------------|--------|-------|-------|-------|
| TEST | N | . GS | AR AR | UK | ., PC | NO | cs | AS | HK | MC | ΕI |
| TIMES | | | | | _ | -0.021 | | | | | |
| TIME2 | 73 | | 0.200 | | | -0.027 | | | | | 0.222 |
| 11MG2 | /3 | 0.346 | 0.171 | 0.220 | 0.217 | -0.027 | 0.240 | 0.207 | 0.20 | 0.501 | U.LLL |
| 43132 | (APPRI | . STRAT | EGIC AI | RCRAFT | MAINTE) | NANCE SP | ECIALIS | (T) | | | |
| TEST | N | GS. | AR | uk. | PC | NO NO | CS | AS | HK | MC | Εľ |
| TIME1 | | | | | | 0.251 | | | | 0.306 | 0.460 |
| TIMEZ | | | | | | 0.348 | | | | | |
| ****** | • | 0.505 | •••• | 0.2.0 | 0.400 | | | | | | |
| 42335 | (APPRN | . AEROS | PACE GR | OUND E | XUI PHENT | MECHAN | IC) | | | | |
| TEST | N | GS | AR | LIK | PC | NO | CS | AS | MK | MC | 13 |
| TIME1 | 71 | 0.090 | 0.446 | -0.035 | 0.061 | 0.375 | 0.026 | 0.263 | 0.268 | 0.232 | 0.207 |
| TIME2 | 71 | | | | | 0.439 | | | | | 0.330 |
| | | | | | | | | | | | |
| 46130 | (APPRN | . MUNIT | IONS SY | STEMS S | PECIAL | (TZ | | | | | |
| TEST | M | GS | AR | UK | PC | NO | CS | AS | HK | NC | Ei |
| TIME1 | 55 | 0.392 | 0.422 | 0.351 | 0.343 | 0.458 | 0.175 | 0.447 | 0.366 | 0.271 | 0.386 |
| TIME2 | 55 | 0.623 | 0.398 | 0.166 | 0.108 | 0.257 | 0.094 | 0.479 | 0.421 | 0.427 | 0.586 |
| | | | | | | | | | | | |
| 64531 | (APPRN | . MATER | HEL STO | RAGE AN | D DISTE | RIBUTION | SPEC1A | LIST) | | | |
| TEST | M | G\$ | AR | WK | PC | NO | cs | AS | MK | MC | El |
| TIME | 52 | 0.261 | 0.254 | 0.383 | 0.349 | 0.211 | 0.372 | 0.087 | 0.161 | 0.149 | 0.034 |
| TIMEZ | 52 | 0.290 | 0.146 | 0.371 | 0.473 | 0.113 | 0.196 | -0.046 | 0.204 | 0.110 | 0.208 |
| | | | | | | | | | | | |
| 12230 | (APPRN | . AIRCE | REW LIFE | SUPPOR | RT SPEC | (ALIST) | | | | | |
| TEST | N | GS. | AR | WK | PC | NO | CS | AS | MK | MC | ΕI |
| TIMET | | | | | | 0.004 | | | | | |
| TIME2 | 52 | 0.278 | 0.175 | 0.251 | 0.264 | 0.090 | 0.235 | 0.259 | 0.262 | 0.281 | 0.270 |
| | | | | | | | | | | | |
| | | | | | | WICE SPE | | | | | |
| TEST | M | GS | AR _ | WK | PC | NO | C\$ | AS. | MK | MC | EI |
| TIMET | 45 | | | | | 0.183 | | | | | |
| TIXME2 | 45 | 0.318 | 0.321 | 0.288 | 0.421 | 0.148 | 0.236 | 0.307 | 0.460 | 0.480 | 0.320 |
| | | | | | | | | | | | |
| | | | | | | IONS SPE | | | | | |
| TEST | M | GS | AR | WK | PC | NO | CS | AS | MK | KC | EI |
| TIME | | | | | | 0.230 | | | | | |
| TIME2 | 43 | 0.519 | 0.174 | 0.586 | 0.403 | 0.288 | 0.189 | 0.342 | 0.330 | 0.432 | 0.370 |
| | | | | | | | | | | | |
| | (APPRN | | | | | IS SPECI | | | | | |
| TEST | W | G\$ | AR | WK | PC | NO | cs | AS | MK | HC | ΕI |
| TIME | | 0.075 | | | | -0.262 | | | 0.233 | | 0.101 |
| TIME2 | 43 | 0.251 | 0.212 | -0.028 | 0.162 | -0.112 | 0.037 | 0.186 | 0.170 | 0.141 | 0.137 |
| | | | | | | | | | | | |
| 60531 | (APPRN | . AIR C | ARGO SP | PECIALIS | | | | | | | |
| TEST | × | G S | AR | SK | PC | MO | c s | AS | MK | MC | 13 |
| TIME 1 | 40 | | | | | -0.043 | | | | | |
| TIME2 | 40 | 0.295 | 0.334 | 0.273 | 0.159 | -0.040 | 0.105 | 0.135 | 0.254 | 0.136 | 0.114 |

Table 5

DIFFERENCES IN TEST VALIDITIES FOR COURSES, ORDERED BY TESTS AND T-VALUES

| | | _ | | | _ | | | _ | ~~ | | _ |
|--------|----|--------------------|--------|------------|--------|--------|----|------------------------|--------|--------------|--------|
| COURSE | | 1 | COURSE | | 1 | COURSE | | T 4 77 0 | COURSE | | T |
| 81132 | AR | -1.308 | 81130 | CS | 0.038 | 90230 | MC | -1.732 | 73230 | NO. | -0.313 |
| 81150 | AR | -1.200 | 70230 | CE | 0.090 | 64530 | HC | -1.704 | 81130 | NO | -0.190 |
| 57130 | AR | -1.082 | 46130 | CS | 0.094 | 57130 | MC | -1.673 | 60531 | NO | 0.046 |
| 73230 | AR | -1.046 | 29130 | CS | 0.112 | 46130 | NC | -1.246 | 43133 | NO | 0.236 |
| 42330 | AR | -0.938 | 81132 | cs | 0.502 | 73230 | NC | -0.973 | 12230 | MO | 0.639 |
| 70230 | AR | -0.105 | 42632 | CS | 0.850 | 43133 | HC | -0.881 | 64531 | NO | 0.807 |
| 90230 | AR | -0.082 | 64530 | cs | 0.896 | 43132 | HC | -0.712 | 43131 | NC | 0.905 |
| 43131 | AR | 0.000 | 57130 | CS | 1.002 | 12230 | MC | -0.497 | 63130 | NO | 0.977 |
| 60531 | AR | 0.097 | 43132 | CS | 1.591 | 60531 | MC | -0.475 | 90230 | NO | 1.147 |
| 42632 | AR | 0.119 | 64531 | cs | 1.823 | 42335 | NC | -0.234 | 46130 | NO | 2.557 |
| 46130 | AR | 0.223 | | | | 63130 | MC | -0.204 | | | |
| 43133 | AR | 0.232 | 64531 | ΕI | -2.054 | 29130 | MC | -0.151 | 29130 | PC | -1.702 |
| 42335 | AR | 0.320 | 46130 | ΕI | -2.008 | 64531 | MC | 0.351 | 70230 | PC | -1.701 |
| 63130 | AR | 0.485 | 90230 | EI | -1.972 | 81132 | MC | 0.378 | 81130 | PC | -1.190 |
| 43132 | AR | 0.485 | 43131 | EI | -1.946 | 81150 | MC | 0.921 | 64531 | PC | -1.091 |
| 81130 | AR | 1.051 | 73230 | ΕI | -1.849 | 43131 | MC | 0.973 | 42335 | PC | -0.801 |
| 64531 | AR | 1,110 | 81150 | ΕI | -1.824 | 81130 | KC | 1.086 | 43133 | PC | -0.768 |
| 12230 | AR | 1.184 | 81132 | ΕI | -1.675 | 42632 | HC | 1.361 | 12230 | PC | -0.752 |
| 29130 | AR | 1.371 | 81130 | EI | -1.497 | 42330 | MC | 1.489 | 73230 | PC | -0.396 |
| 64530 | AR | 1.874 | 42335 | EI | -1.305 | 70230 | MC | 1. <i>7</i> 22 | 43132 | PC | -0.287 |
| | | | 64530 | Ei | -1.218 | | | | 63130 | PC | -0.231 |
| 81132 | AS | -1.887 | 63130 | ΕI | -1.097 | 42335 | MK | -2.421 | 81150 | PC | -0.068 |
| 90230 | AS | -1.823 | 57130 | EI | -0.628 | 70230 | MK | -1.928 | 90230 | PC | 0.194 |
| 60531 | AS | -1.787 | 12230 | E1 | -0.531 | 43133 | HK | -0.898 | 64530 | PC | 0.217 |
| 70230 | AS | -1.644 | 42330 | El | -0.271 | 64530 | K | -0.865 | 60531 | PC | 0.321 |
| 43131 | AS | ~1.57 9 | 60531 | EI | 0.000 | 46130 | MK | -0.621 | 43131 | PC | 0.740 |
| 43133 | AS | -0.942 | 43133 | El | 0.384 | 64531 | MK | -0.558 | 42632 | PC | 1.009 |
| 43132 | AS | -0.376 | 29130 | EI | 0.417 | 43131 | HK | -0.545 | 42330 | PC | 1.143 |
| 46130 | AS | -0.356 | 43132 | Εi | 0.795 | 43132 | MK | -0.500 | 57130 | PC | 1.158 |
| 81150 | AS | -0.260 | 42632 | ΕI | 0.810 | 29130 | MK | -0.135 | 81132 | PC | 1.304 |
| 64530 | AS | -0.121 | 70230 | EI | 1.131 | 12230 | MK | -0.082 | 46130 | PC | 1.609 |
| 81130 | AS | -0.094 | | | | 42632 | MK | -0.037 | | | |
| 57130 | AS | -0.075 | 81130 | GS | -2.717 | 60531 | MK | 0.120 | 29130 | WK. | -2.436 |
| 42330 | AS | 0.094 | 46130 | GS | -2.619 | 81150 | MK | 0.271 | 12230 | WK | -1.128 |
| 42335 | AS | 0.095 | 12230 | C2 | -1.699 | 63130 | MK | 0.447 | 90230 | WK | -1.125 |
| 73230 | AS | 0.279 | 81132 | GS | -1.630 | 81130 | MK | 0.511 | 42335 | WK. | -0.869 |
| 29130 | AS | 0.415 | 57130 | GS | -1.628 | 73230 | MK | 0.541 | 60531 | u K | -0.835 |
| 42632 | AS | 0.769 | 43132 | CS | -1.500 | 42330 | MK | 0.652 | 42632 | LIK | -0.238 |
| 63130 | AS | 1.491 | 29130 | G\$ | -1.390 | 81132 | MK | 0.762 | 42330 | WK | -0.213 |
| 64531 | AS | 1.582 | 42330 | G\$ | -1.245 | 90230 | MK | 1.486 | 43132 | WK | 0.082 |
| 12230 | AS | 2.091 | 60531 | CS | -1.183 | 57130 | MK | 2.059 | 63130 | i a K | 0.094 |
| | | | 42335 | GS | -0.994 | | | | 64531 | WK | 0.112 |
| 42335 | CS | -1.927 | 42632 | GS | -0.831 | 81132 | NO | -1.305 | 57130 | WK | 0.125 |
| 73230 | CS | -1.728 | 81150 | G\$ | -0.416 | 70230 | NO | -1.235 | 81150 | WK | 0.161 |
| 81150 | CS | -1.231 | 70230 | GS | -0.347 | 42330 | NO | -1.147 | 43133 | WK | 0.251 |
| 90230 | cs | -1.028 | 64531 | GS | -0.294 | 43132 | NO | -0.970 | 81130 | UK | 0.547 |
| 42330 | CS | -0.470 | 90230 | CS | 0.130 | 81150 | MO | -0.894 | 81132 | ₩ | 0.863 |
| 63130 | cs | -0.442 | 64530 | G S | 0.230 | 42335 | NO | -0.638 | 73230 | WK | 0.955 |
| 43131 | ÇS | -0.291 | 43133 | G\$ | 0.663 | 64530 | NO | -0.605 | 64530 | WK | 0.993 |
| 12230 | cs | -0.220 | 43131 | GS | 0.918 | 57130 | NO | -0.598 | 43131 | ٧K | 1.519 |
| 43133 | C2 | -0.172 | 63130 | GS | 0.977 | 42632 | NO | -0.550 | 46130 | WK | 1.608 |
| 60531 | CS | -0.069 | 73230 | GS | 1.200 | 29130 | NO | -0.546 | 70230 | WK | 1.884 |
| | | | | | | | | | | | |

Table 6

META-ANALYSES RESULTS

20 SCHOOLS

| TEST GROUP | T1 | T2 | R12 | D <i>F</i> | T |
|------------------|-------|-----------|-------|------------|--------|
| Tech. Knowledge | 0.286 | 0.322 | 0.737 | 7321 | -4.499 |
| General Ability | 0.336 | 0.335 | 0.727 | 7321 | 0.124 |
| Perceptual Speed | 0.141 | 0.148 | 0.668 | 3659 | -0.526 |
| Total | 0.280 | 0.294 | 0.720 | 18307 | -2.662 |
| | | 69 SCHOOL | .s | | |
| TEST GROUP | 71 | T2 | R12 | DF | T |
| Tech. Knowledge | 0.295 | 0.321 | 0.737 | 11337 | -4.047 |
| General Ability | 0.315 | 0.315 | 0.727 | 11337 | 0.000 |
| Perceptual Speed | 0.124 | 0.136 | 0.668 | 5667 | -1.120 |
| Total | 0.270 | 0.283 | 0.720 | 28347 | -3.065 |

Note: Values under T1 and T2 are average validity coefficients computed using Fisher's z-transformation and weighted by N's. The R12 column is the average time 1-time 2 test intercorrelations, again computed using Fisher's z's and weighted by N's. The last column reports Hotelling's t's based on correlated data for the differences in test and retest average validities within each category.

in Table 6, not only for data in the 20 schools reported in Table 5, but for all schools in which the matching process yielded 15 or more subjects. The general approach was to treat data from each school as being from a separate study. The data were combined using a method recommended by Hedges and Olkin (1985, pp. 230-232). First, the test and retest validities within each category were converted to Fisher's z's and weighted by N-3. Next, the weighted z's were averaged, and the averages were converted back into correlations. Finally, Hotelling's t's were computed for the differences in test and retest average correlations within each subtest category.

The data in Table 6 reveal that there are no significant differences between Time 1 (test) and Time 2 (retest) validities for the ASVAB subtests associated with either general learning ability or

perceptual speed. Subtest validities for the technical knowledge area are significantly higher at Time 2, both in the 20-school and 69-school samples, but the magnitudes of the differences are very small. The average retest validities across all schools in both the 20-school and 69-school samples are significantly higher at Time 2, but this was in the presence of many degrees of freedom, and due solely to differences in the technical knowledge subtests.²

In summary, analyses of the zero-order validities present a clear indication that ability changes for Air Force applicants between time of testing and entry into the service are restricted to measures of technical knowledge, and are relatively minor.

² Some observers may object to the meta-analysis approach because subjects attending the various schools cannot be assumed to have been drawn from the same population. This is certainly 'rue, since the entry ability level requirements vary from course to course. In order to respond to this possible criticism, all validity values in the 20-school sample were corrected for restriction in range due to selection, and the average validities were recomputed. The revised Time 1-Time2 average validities were: technical knowledge, .405 vs. .425.; general ability, .501 vs. .485; perceptual speed, .197 vs .215, and total, .408 vs. .411. These differences appear to be less than those computed for the uncorrected validities. Hotelling t's were not computed because of problems in dealing with double curtailments associated with the intercorrelation terms.

Regression Analyses. Another way of evaluating the impact of ability changes is to determine the differences in these multiple correlations against technical school grades which are obtained using Time 1 and Time 2 data. This was accomplished for each of the 20 technical schools courses in which the matching process yielded 40 or more subjects. Again, the mean multiple R's were based upon weighted z's. The results yielded a multiple R of .546 for Time 1 (test) and .539 for Time 2 (retest). Thus, even considering the joint action of all ASVAB subtests, there is no indication of appreciable loss of predictive power across time.

Internal Regression Analyses Using ASVAB Test-Retest Data

It appears from the preceding analyses that, if ability changes occur for Air Force selectees between time of operational testing and time of entry into the service, the changes are associated with technical knowledge and are relatively minor. Estimates of the value of new tests in adding power to the operational ASVAB in predicting technical school grades would appear to be fairly trustworthy. However, on many occasions LAMP experi-

mental tests are validated against intermediate learning criteria which are administered in the same experimental testing session. It is entirely possible that the correlations between the experimental tests and criterion measures are inflated because of test covariance associated with motivational decrements present in an experimental testing situation. It has been conjectured that some individuals operate with decreased motivation when they are told that the test results will be used only for purposes of research, and this negatively impacts their performance on both experimental and criterion measures administered during an experimental testing session. This type of behavior would be expected to increase the correlations between the experimental and criterion measures. In the case of LAMP validation studies, where the contribution of experimental tests to ASVAB is being evaluated against a concurrent criterion measure, it is important to know how much of the computed contribution might be associated with this type of situational variance.

Data indicating existence of the problem are shown in Table 7. The last three columns of this table report the average correlation (computed using Fisher's z-transformations) of each ASVAB

TABLE 7

TIME1,TIME2 MEANS AND AVERAGE INTERCORRELATIONS FOR TIME1,TIME2 AND TIME1-TIME2

TOTAL SAMPLE (N=4077)

| | MEAN SC | ORES | STD.DE | VIATIONS | AVG.INTERCORRELATIONS | | | |
|------|---------|-------|--------|----------|-----------------------|-------|-------|--|
| TEST | TIME1 | TIME2 | TIME1 | TIME2 | T1-T1 | T1-T2 | T2-T2 | |
| GS | 18.66 | 18.59 | 3.56 | 3.71 | 0.375 | 0.377 | 0.393 | |
| AR | 22.89 | 22.08 | 4.60 | 5.22 | 0.364 | 0.377 | 0.401 | |
| WK | 29.05 | 29.50 | 4.33 | 4.16 | 0.322 | 0.324 | 0.357 | |
| PC | 12.44 | 12.37 | 1.91 | 1.95 | 0.285 | 0.278 | 0.314 | |
| NO | 40.91 | 40.52 | 7.12 | 7.59 | 0.131 | 0.144 | 0.193 | |
| CS | 52.83 | 53.70 | 11.67 | 12.18 | 0.133 | 0.130 | 0.159 | |
| AS | 18.25 | 18.62 | 4.37 | 4.34 | 0.241 | 0.232 | 0.267 | |
| MK | 16.81 | 16.31 | 4.87 | 5.16 | 0.368 | 0.365 | 0.392 | |
| MC | 17.74 | 17.96 | 4.03 | 3.97 | 0.349 | 0.343 | 0.350 | |
| ΕI | 14.00 | 13.94 | 3.14 | 3.38 | 0.320 | 0.300 | 0.359 | |
| AVG. | 24.36 | 24.36 | 4.96 | 5.17 | 0.291 | 0.290 | 0.320 | |

subtest with the other nine subtests, separately for Time 1, for Time 2, and between Time 1 and Time 2. Notice that the Time 1-Time 2 correlations are almost equivalent to the Time 1-Time 1 correlations, providing additional evidence on the lack of ability changes across time. However, notice that the Time 2-Time 2 correlations are systematically higher than either the Time 1-Time 1 or the Time 1-Time 2 correlations. This may be attributable to common situational variance associated with the experimental testing situation at Time 2. That is, some individuals were either tired, under stress, or unmotivated during the Time 2 experimental retesting session, and they systematically made low scores on all subtests. One would expect this factor to result in lower means, but comparing the means from Time 1 to those from Time 2 is meaningless because of the interacting factors of learning, forgetting, practice, and regression effects mentioned previously.

The standard deviations appear to be generally higher at Time 2, and this too is attributed to situational variance present in the experimental testing session. Poorly motivated subjects could actually score so low that they would not have been accepted into the Air Force had they performed at the same level during the operational administration.

The fact that situational variance in an experimental testing session increases test intercorrelations has implications for the LAMP program. It indicates that the correlations observed between experimental tests and criterion tests administered in the same experimental testing session are likely to be inflated. Furthermore, the computed contribution of the experimental tests to ASVAB scores collected some months previously would also be inflated. The question is how seriously such estimates are inflated.

The ASVAB test-retest data were used to roughly evaluate the amount of inflation likely in estimating the contribution of experimental tests to ASVAB. The approach was to use each of eight ASVAB subtest scores at Time 2 as a criterion measure. In each instance, a related subtest was

selected as being an experimental predictor variable, while the remaining eight ASVAB subtest scores at Time 1 and again at Time 2 were treated as though they were the complete ASVAB. Three comparisons were evaluated for each of the eight criterion measures as defined by the experimental design shown in Table 8.

Table 9 reports the R-squares for each of the six comparison equations, separately for eight criteria. The NO and CS subtests were not used as criteria in this analysis because of their unreliability, and because they tend to correlate only with each other.

Table 10 summarizes differences between Comparisons 1, 2, and 3 from Table 9. Comparison 1 is the estimated contribution of an experimental test to ASVAB, when ASVAB is administered at Time 1, and data for the experimental test and criterion data are collected at Time 2 in the same experimental testing session. Comparison 2 reports the actual contribution of the experimental test to ASVAB. In this latter instance, the experimental test and ASVAB data are collected at Time 1, and the criterion data are collected at Time 2. The difference between Comparison 1 and Comparison 2 estimates reflects the degree of inflation in the estimated contribution of an experimental test to ASVAB, when both the experimental test and criterion data are collected concurrently in the same experimental testing session. It appears that such inflation is not very serious. As shown in Table 10, the average estimated variance contribution across all eight criteria in the total sample is 12.8%, while the actual contribution is 10.4%. Application of an 18.75% discount would bring the average estimated contribution into alignment.

The data on Comparison 3 presented in Table 10 suggest that if ASVAB were readministered along with the experimental tests and criterion measures, the estimated contribution of the experimental tests to ASVAB would generally not be inflated. If sufficient testing time were available, this might be the preferred approach.

Table 8

EXPERIMENTAL DESIGN

Y= CRITERION (ONE OF EIGHT ASVAB TESTS).

X= PREDICTOR TEST BEING EVALUATED (A RELATED ASVAB TEST).

ASVAB= REMAINING 8 TESTS IN ASVAB

TIME1= TIME OF OPERATIONAL ADMINISTRATION OF ASVAB

TIME2= 6TH DAY OF BASIC TRAINING. SAME DAY AS LAMP TESTS ARE ADMINISTERED.

COMPARISON 1

R-SQUARED1

Y(TIME2) AS CRITERION

X(TIME2) AND ASVAB(TIME1) AS PREDICTORS

R-SQUARED2

Y(TIME2) AS CRITERION

ASVAB(TIME1) AS PREDICTORS

(R-SQUARED1) - (R-SQUARED2) STIMATED UNIQUE CONTRIBUTION OF EXPERIMENTAL TEST TO ASVAB IN ACCOUNTING FOR CRITERION VARIANCE. THIS IS THE WAY THE AIR FORCE AND OTHER SERVICES ESTIMATE THE VALUE OF NEW TESTS. THIS ESTIMATE IS KNOWN TO BE BIASED IN AN UPWARD DIRECTION, BUT HOW MUCH IS UNKNOWN.

COMPARISON 2

R-SQUARED3

Y(TIME2) AS CRITERION

X(TIME1) AND ASVAB(TIME1) AS PREDICTORS

R-SQUARED4

Y(TIME2) AS CRITERION

ASVAB(TIME1) AS PREDICTORS

(R-SQUARED3) - (R-SQUARED4)= <u>ACTUAL</u> UNIQUE VARIANCE CONTRIBUTED BY EXPERIMENTAL TEST TO ASVAB IN ACCOUNTING FOR CRITERION VARIANCE, WHEN THE EXPERIMENTAL TEST IS ADMINISTERED ON THE SAME OCCASION AS THE OPERATIONAL ASVAB.

COMPARISON 3

R-SQUAREDS

Y(TIME2) AS CRITERION

X(TIME2) AND ASVAB(TIME2) AS PREDICTORS

R-SQUARED6

Y(TIME2) AS CRITERION

ASVAB(TIMEZ) AS PREDICTORS

(R-SQUARED5) - (R-SQUARED6)= UNIQUE VARIANCE CONTRIBUTED BY THE EXPERIMENTAL TEST TO ASVAB IN ACCOUNTING FOR THE CRITERION VARIANCE, WHEN ASVAB IS READMINISTERED CONCURRENTLY WITH THE CRITERION AND EXPERIMENTAL TESTS.

Table 9

SQUARED MULTIPLE CORRELATIONS FOR COMPARISON EQUATIONS

| | TOTAL SAMPLE (N=4077) | | | | | 0-6MOS SAMPLE (N=1774) | | | | |
|--------|-----------------------|-----------|----------|-------|-------|------------------------|-----------|--------|---------|---------|
| TESTS | EQU1 | EQU2&4 | EQU3 | EQU5 | EQU6 | EQU1 | EQU2&4 | EQU3 | EQU5 | EQU6 |
| | R-SQ | R-SQ | R-SQ | R-SQ | R-SQ | R-SQ | R-SQ | R-SQ | R-SQ | R-SQ |
| C - P | | | | | | | | | | |
| GS-WK | 0.511 | 0.384 | 0.503 | 0.536 | 0.442 | 0.511 | 0.398 | 0.510 | 0.531 | 0.448 |
| AR-MK | 0.527 | 0.342 | 0.487 | 0.550 | 0.390 | 0.549 | 0.362 | 0.498 | 0.573 | 0.416 |
| WK-GS | 0.464 | 0.326 | 0.431 | 0.486 | 0.383 | 0.477 | 0.357 | 0.456 | 0.504 | 0.417 |
| PC-WK | 0.318 | 0.199 | 0.275 | 0.342 | 0.260 | 0.355 | 0.237 | 0.310 | 0.376 | 0.283 |
| AS-HC | 0.468 | 0.382 | 0.445 | 0.486 | 0.418 | 0.463 | 0.373 | 0.436 | 0.482 | 0.417 |
| KX-AR | 0.556 | 0.383 | 0.548 | 0.550 | 0.398 | 0.573 | 0.374 | 0.550 | C.539 | 0.415 |
| MC-AS | 0.453 | 0.364 | 0.426 | 0.471 | 0.401 | 0.448 | 0.357 | 0.426 | 0.471 | 0.404 |
| EI-AS | 0.528 | 0.422 | 0.517 | 0.531 | 0.431 | 0.518 | 0.412 | 0.516 | 0.526 | 0.430 |
| | 6-1 | IZHOS SAI | MPLE (N= | 1785) | | ď | VER 12MOS | SAMPLE | (N=518) | |
| TESTS | EQU1 | EQU2&4 | EQU3 | EQU5 | EQU6 | EQU1 | EQU2&4 | EQU3 | EQU5 | EQU6 |
| | R-SQ | R-SQ | R-SQ | R-SQ | R-SQ | R-SQ | R-SQ | R-SQ | R - SQ | R - \$0 |
| C - P | | | | | | | | | | |
| GS-WK | 0.513 | 0.376 | 0.506 | 0.535 | 0.430 | 0.510 | 0.381 | 0.496 | 0.564 | 0.467 |
| AR-HK | 0.516 | 0.340 | 0.481 | 0.540 | 0.379 | 0.479 | 0.276 | 0.480 | 0.503 | 0.329 |
| WK-G\$ | 0.451 | 0.301 | 0.410 | 0.474 | 0.356 | 0.468 | 0.326 | 0.436 | 0.465 | 0.357 |
| PC-WK | 0.279 | 0.158 | 0.245 | 0.301 | 0.224 | 0.324 | 0.221 | 0.277 | 0.384 | 0.329 |
| AS-MC | 0.478 | 0.394 | 0.457 | 0.492 | 0.421 | 0.463 | 0.386 | 0.438 | 0.495 | 0.434 |
| MK-AR | 0.551 | 0.385 | 0.537 | 0.542 | 0.398 | 0.525 | 0.322 | 0.548 | 0.506 | 0.331 |
| MC-AS | 0.452 | 0.364 | 0.423 | 0.474 | 0.401 | 0.472 | 0.392 | 0.442 | 0.474 | 0.416 |
| EI-AS | 0.528 | 0.426 | 0.509 | 0.526 | 0.417 | 0.573 | 0.470 | 0.564 | 0.578 | 0.486 |

NOTES: Equations are numbered as they are described in the experimental design.

Table 10

ESTIMATED CONTRIBUTION OF EXPERIMENTAL TESTS TO ASVAB IN THE PREDICTION OF SELECTED CRITERION MEASURES

| | | TOTAL S | AMPLE (N=4 | ,077) | 0-6MOS SAMPLE (N=1,774) | | | |
|-----------|-----------|--------------------------|------------|--------------------------|-------------------------|---------|---------|--|
| | | COMP. 1 | COMP. 2 | COMP. 3 | COMP. 1 | COMP. 2 | COMP.3 | |
| CRI | PRE | EQ1-EQ2 | E93-E92 | E95-E94 | EQ1-EQ2 | EQ3-EQ2 | EQ5-EQ4 | |
| GS | ₩K | 12.7 | 11.9 | 9.4 | 11.3 | 11.2 | 8.3 | |
| AR | HK | 18.5 | 14.5 | 16.0 | 18.7 | 13.6 | 15.7 | |
| WK | GS | 13.8 | 10.5 | 10.3 | 12.0 | 9.9 | 8.7 | |
| PC | WK | 11.9 | 7.6 | 8.2 | 11.8 | 7.3 | 9.3 | |
| AS | MC | 8.6 | 6.3 | 6.8 | 9.0 | 6.3 | 6.5 | |
| MK | AR | 17.3 | 16.5 | 15.2 | 17.9 | 16.5 | 15.4 | |
| MC | AS | 8.9 | 6.2 | 7.0 | 9.1 | 6.9 | 6.7 | |
| ΕI | AS | 10.6 | 9.5 | 10.0 | 10.6 | 10.4 | 9.6 | |
| j | KEAN | 12.8 | 10.4 | 10.4 | 12.6 | 10.3 | 10.0 | |
| | | 6-12MOS SAMPLE (N=1,785) | | OVER 12MOS SAMPLE (N=518 | | | | |
| | | COMP. 1 | COMP. 2 | COMP. 3 | COMP. 1 | COMP. 2 | COMP. 3 | |
| CRI | PRE | EQ1-EQ2 | EQ3-EQ2 | EQ5-EQ4 | EQ1-EQ2 | EQ3-EQ2 | EQ5-EQ4 | |
| GS | WK | 13.7 | 13.0 | 10.5 | 12.9 | 11.5 | 9.7 | |
| AR | HK | 17.6 | 14.1 | 16.1 | 20.3 | 20.4 | 17.4 | |
| WK | GS | 15.0 | 10.9 | 11.8 | 14.2 | 11.0 | 10.8 | |
| PC | WK | 12.1 | 8.7 | 7.7 | 10.3 | 5.0 | 5.5 | |
| AS | HC | 8.4 | 6.3 | 7.1 | 7.7 | 5.2 | 6.1 | |
| MK | AR | 16.6 | 15.2 | 14.4 | 20.3 | 22.6 | 17.5 | |
| HC | AS | 8.8 | 5.9 | 7.3 | 8.0 | 5.0 | 5.8 | |
| ΕI | AS | 10.2 | 8.3 | 10.9 | 10.3 | 9.4 | 9.2 | |
| | 4EAN | 12.8 | 10.3 | 10.7 | 13.0 | 11.3 | 10.3 | |

NOTES: Reported values are differences in R-squares. Equations are numbered as they are described in the experimental design.

COMP. 1= CRITERION & EXPERIMENTAL TEST AT TIME 2, ASVAB AT TIME 1
Criterion & experimental test at Time 2, ASVAB at Time 1.

COMP. 2= CRITERION AT TIME 2, EXPERIMENTAL TEST AND ASVAB AT TIME 1
Criterion at Time 2, experimental test and ASVAB at Time 1.

COMP. 3= CRITERION, EXPERIMENTAL TEST, & ASVAB AT TIME2
Criterion, experimental test & ASVAB at Time 2.

V. DISCUSSION

The basic problem addressed in this paper is how to estimate the contribution of an experimental test to ASVAB in predicting a criterion measure. The only sure method would be to administer the experimental test along with the ASVAB at official testing stations under operational conditions. Then one could estimate the contribution of the experimental test by evaluating the difference in R-squares yielded by a full model (containing the ASVAB subtests and the experimental variable as predictors) and a restricted model (using only ASVAB subtests as predictors). Because a large amount of additional testing time in the operational environment would be required by this approach, it is usually not feasible.

In the typical situation, experimental tests are administered to individuals already in the service and an attempt is made to see how much predictive efficiency is added by these experimental tests to ASVAB subtests administered some months previously. This approach has been criticized, in that ability levels could have changed between the time of ASVAB testing and experimental test administration. The present study is an attempt to evaluate the changes in the abilities of Air Force personnel between the time of their selection and their entry into service.

With the exception of slightly higher retest validities for ASVAB measures of technical knowledge, the test and retest validities centered at about the same level. Although the validities of Time 2 ASVAB scores may have been slightly underestimated due to variance associated with a reduced level of motivation for some examinees during the experimental testing session, this same factor will lead to underestimates of the validity of any experimental test for subsequent operational criteria. It is concluded that reasonable trust can

be placed in estimates of the contributions of experimental tests to ASVAB in the prediction of school grades, even though the ASVAB scores are collected some months prior to administration of the experimental tests.

Results of this study also address a potentially more serious problem. Experimental test scores and intermediate learning criterion scores are often collected in the same experimental testing session. In this case, the correlation computed between experimental and criterion test scores can be inflated by variance associated with the experimental testing situation, and the estimated contributions of the experimental tests to ASVAB are likely to be inflated. Results from analyses conducted using the ASVAB test-retest data indicate that such inflation does indeed exist, but it is not large. The average estimated contribution of an experimental test to ASVAB across eight criterion measures was 12.8%, while the actual computed contribution when the experimental tests were administered along with ASVAB at Time 1 was 10.4%.

Another approach to estimating the contribution of an experimental test to ASVAB would be to readminister ASVAB concurrently with the experimental test. Data in this study indicate that this approach would, on the average, yield values comparable to those which would be obtained if the experimental test had been administered at the testing stations, along with ASVAB.

VI. CONCLUSIONS

With regard to validation studies against concurrent criteria, some inflation in the estimated contribution of experimental tests to ASVAB should be expected. This appears to be due to a subset of individuals who, because of factors occurring in an

experimental testing situation, do poorly on both experimental tests and concurrently administered criterion measures. Even if no adjustments are made, the degree of inflation in the estimated contributions is not large. Results from the present investigation suggest that an 18.8% reduction in the estimated contribution would bring such estimates closer to a stual contributions. In many studies, such discounts may not be necessary, if an attempt is made to identify and remove unmotivated subjects from analyses.³

Until better information becomes available, it is recommended that the estimated contribution of experimental tests to ASVAB in the prediction of subsequent criteria be accepted at face value. In the case of concurrent criteria, it is recommended that an attempt be made to eliminate obviously unmotivated subjects from analyses (such as those who score at or near chance levels on all variables). Estimated contributions may still be slightly inflated, but present results suggest that such inflation is not likely to be large.

References

Hedges, L.V., & Olkin, I. (1985). statistical methods for meta-analysis. Orlando: Academic Press, Inc.

Humphreys, L.G. (1960). Investigations of the simplex. Psychometrika, 25, 475-483

Humphreys, L.G., & Davey, T.C. (1988). Continuity in intellectual growth from 12 months to 9 years. Intelligence, 12, 183-197.

Humphreys, L.G., Davey, T.C, & Park, R.K. (1985). Longitudinal correlation analysis of standing height and intelligence. *Child Development*, <u>56</u>, 1465-1478.

Humphreys, L.G., Parsons, C.K., & Park, R.K. (1979). Application of a simplex process model to six years of cognitive development in four demographic groups. Applied Psychological Measurement, 3, 51-64.

Ree, J.M., Mullins, C.J., Mathews, J.J., & Massey, R.H. (1982). Armed Services Vocational Aptitude Battery: Item and factor analyses of Forms 8, 9, and 10. (AFHRL-TR-81-55, AD-A113 465). Brooks AFB, TX: Manpower and Personnel Division, Air Force Human Resources Laboratory.

³ There is a special problem in evaluating tests of information processing speed. No comparable tests were evaluated in the present study, and there is reason to believe that highly speeded tests are the ones most likely to suffer from situational variance during experimental testing sessions. It is especially important to remove unmotivated subjects from validation studies involving tests of processing speed.